

Lecture 4

Statistical Inference

with

Info-Gaps in Underlying Processes

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1 “Temperament” Measurement: Testing a Sample Mean

1.1 The Question

§ We wish to sample a population, to estimate T :
Temperament, **T**emperature, emotional **T**urbulence, etc.

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§ How do we **use these measurements to decide**
if the population “really” has temperament T ?

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§ How do we **use these measurements to decide**
if the population “really” has temperament T ?

§ What does “really has temperament T ” *mean?*

§ Perhaps: $E(T) = \mu$. (The mean is the meaning???)

1.2 *Statistical Formulation*

§ Random sample, definition:

- **Intuitively:**

Set of independent measurements from the same population.



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- **Intuitively:**

Set of independent measurements from the same population.

- **More technically:**

Set of independent and identically distributed (i.i.d.) random variables.

§ The **sample mean** is defined as:

$$\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i \quad (1)$$

1.3 *First Hypothesis Test: Normal Distribution*

§ **Suppose:**

- **The temperament is normally distributed.**

or

-

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- The number of measurements is large.

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- The temperament is normally distributed.

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- The number of measurements is large.

§ Thus:

$$\bar{x} \sim \mathcal{N}(\mu, \sigma^2/N) \quad (2)$$

where:

μ = true mean temperament.

σ^2 = variance of the temperament measurements.

N = sample size.

§

§ **Suppose:**

- The temperament is normally distributed.

or

- The number of measurements is large.

§ **Thus:**

$$\bar{x} \sim \mathcal{N}(\mu, \sigma^2/N) \quad (3)$$

where:

μ = true mean temperament.

σ^2 = variance of the temperament measurements.

N = sample size.

§ Also, assume that we **know the value of σ^2** .

§ How do we decide if the true temperament is T ?

§

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§ We use an hypothesis test.

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§ **How do we decide** if the true temperament is T ?

§ We use an **hypothesis test**.

§ The **null hypothesis**:

$$H_0 : \quad \mu = T \quad (4)$$

T = hypothesized temperament: a known value.

μ = true temperament: an unknown value.

§

§ **How do we decide** if the true temperament is T ?

§ We use an **hypothesis test**.

§ The **null hypothesis**:

$$H_0 : \quad \mu = T \quad (5)$$

T = hypothesized temperament: a known value.

μ = true temperament: an unknown value.

§ The **alternative hypothesis**:

$$H_1 : \quad \mu \neq T \quad (6)$$

This is a **two-tailed test**.

§

§ **How do we decide** if the true temperament is T ?

§ We use an **hypothesis test**.

§ The **null hypothesis**:

$$H_0 : \quad \mu = T \quad (7)$$

T = hypothesized temperament: a known value.

μ = true temperament: an unknown value.

§ The **alternative hypothesis**:

$$H_1 : \quad \mu \neq T \quad (8)$$

This is a **two-tailed test**.

§ For a **one-tailed** the **alternative hypothesis** would be:

$$H_1 : \quad \mu > T \quad (9)$$

or:

$$H_1 : \quad \mu < T \quad (10)$$

§ Level of significance, α :

- Probability of obtaining a result at least as extreme as the observed result, conditioned upon H_0 .
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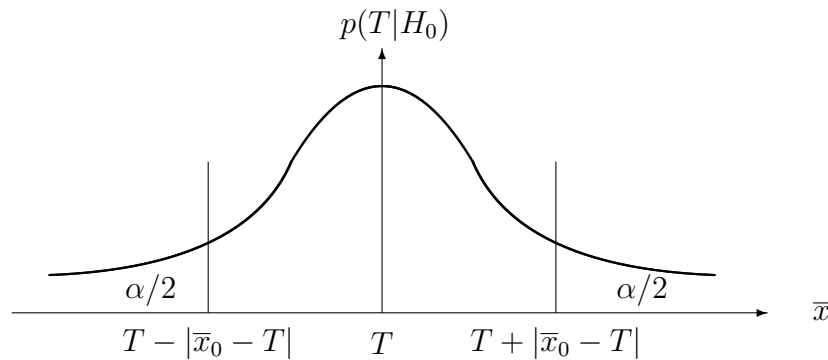


Figure 1: Sketch of probability density, illustrating the level of significance in eq.(11).

§ For the two-tailed test (fig. 1):

$$\alpha = \mathbf{Prob} \left(|\bar{x} - T| \geq |\bar{x}_o - T| \mid H_0 \right) \quad (11)$$

\bar{x} = the random variable “sample mean”.

\bar{x}_o = the observed value of the random variable \bar{x} .

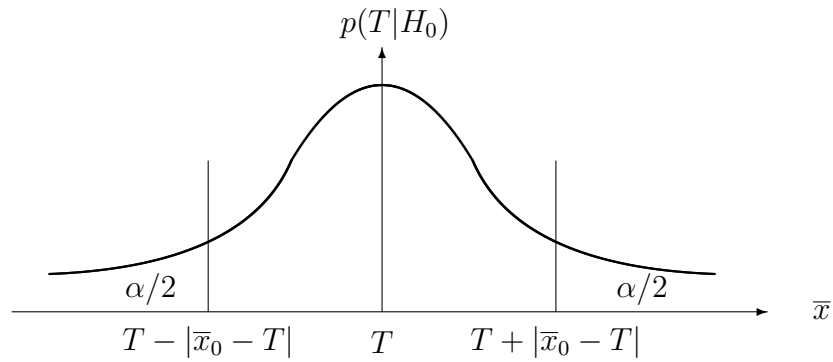


Figure 2: Sketch of probability density, illustrating the level of significance in eq.(11).

§ Interpreting the level of significance:

If α is small: reject H_0 .

If α is large: accept H_0 .

§ How to evaluate the level of significance?

§

§ How to evaluate the level of significance?

§ Conditioned upon H_0 , we can assert

(recall we assume normal measurements or large sample size):

$$\bar{x} \sim \mathcal{N}(T, \sigma^2/N) \quad (12)$$

\bar{x} can be standardized as:

$$z = \frac{\bar{x} - T}{\sigma/\sqrt{N}} \sim \mathcal{N}(0, 1) \quad (13)$$

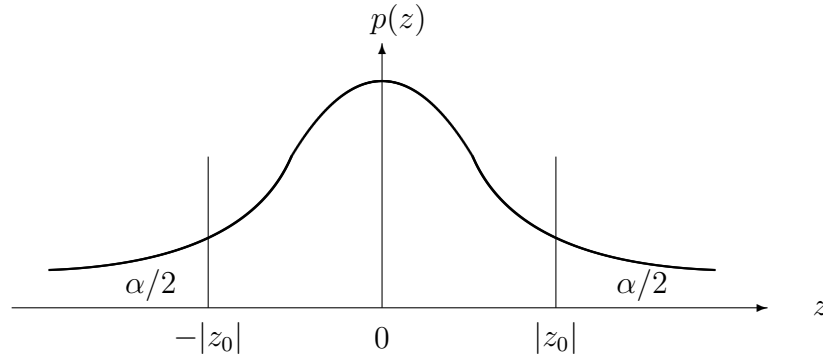


Figure 3: Sketch of probability density illustrating level of significance in eq.(16).

§ **The level of significance can be expressed as (fig. 3):**

$$\alpha = \mathbf{Prob} \left(|\bar{x} - T| \geq |\bar{x}_o - T| \mid H_0 \right) \quad (14)$$

$$= \mathbf{Prob} \left(\frac{|\bar{x} - T|}{\sigma/\sqrt{N}} \geq \frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \mid H_0 \right) \quad (15)$$

$$= 2 \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (16)$$

$\Phi(\cdot)$ is the cumulative distribution function (cdf) of the standard normal distribution.

•

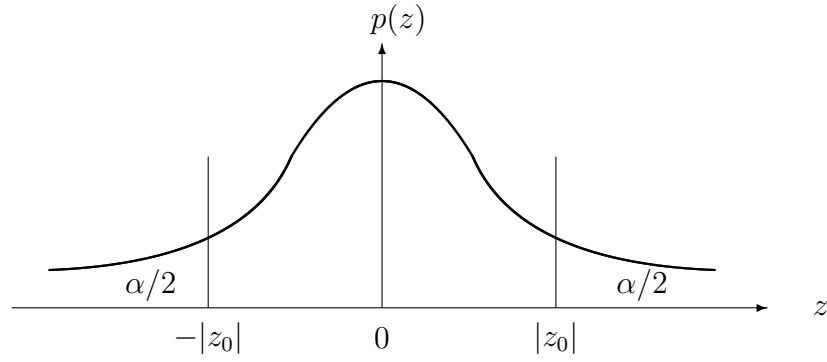


Figure 4: Sketch of probability density illustrating level of significance in eq.(19).

§ **The level of significance can be expressed as (fig. 4):**

$$\alpha = \mathbf{Prob} \left(|\bar{x} - T| \geq |\bar{x}_o - T| \mid H_0 \right) \quad (17)$$

$$= \mathbf{Prob} \left(\frac{|\bar{x} - T|}{\sigma/\sqrt{N}} \geq \frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \mid H_0 \right) \quad (18)$$

$$= 2 \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (19)$$

$\Phi(\cdot)$ is the cumulative distribution function (cdf) of the standard normal distribution.

• **Thus:**

$$z_o = \frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \implies \alpha = 2[1 - \Phi(z_o)] \quad (20)$$

1.4 *Info-Gap Uncertain Distribution Function*

§ The problem.

- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (21)$$

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- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (22)$$

- Assumes the observations are all normal or i.i.d.
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§ The problem.

- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (23)$$

- Assumes the observations are all normal or i.i.d.
- These assumptions may be **wrong** due to:
 - Population heterogeneity.
 -

§ The problem.

- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (24)$$

- Assumes the observations are all normal or i.i.d.
- These assumptions may be wrong due to:
 - Population heterogeneity.
 - Population change during the sampling.
 -

§ The problem.

- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (25)$$

- Assumes the observations are all normal or i.i.d.
- These assumptions may be wrong due to:
 - Population heterogeneity.
 - Population change during the sampling.
 - Non-random sampling procedures.
-

§ The problem.

- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (26)$$

- Assumes the observations are all normal or i.i.d.
- These assumptions may be **wrong** due to:
 - Population heterogeneity.
 - Population change during the sampling.
 - Non-random sampling procedures.
- Hence the cdf, $\Phi(z)$, may be **wrong**.

§

§ The problem.

- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (27)$$

- Assumes the observations are all normal or i.i.d.
- These assumptions may be **wrong** due to:
 - Population heterogeneity.
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 - Non-random sampling procedures.
- Hence the cdf, $\Phi(z)$, may be **wrong**.

§ Uncertainty:

- Extent of error of $\Phi(z)$ unknown.
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§ The problem.

- The level of significance in eq.(16):

$$\alpha = \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (28)$$

- Assumes the observations are all normal or i.i.d.
- These assumptions may be **wrong** due to:
 - Population heterogeneity.
 - Population change during the sampling.
 - Non-random sampling procedures.
- Hence the cdf, $\Phi(z)$, may be **wrong**.

§ Uncertainty:

- Extent of error of $\Phi(z)$ unknown.
- No probabilistic information about this uncertainty.
- This is an **info-gap**.

§ Info-gap model of uncertainty.

- $F(z) =$ unknown correct cdf.
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- Fractional deviation between $F(z)$ and $\Phi(z)$ **unknown**.
- Info-gap model of non-probabilistic uncertainty:

$$\mathcal{U}(h) = \left\{ F(z) : \begin{array}{l} F(-\infty) = 0, \quad F(\infty) = 1, \quad \frac{\partial F(z)}{\partial z} \geq 0, \\ \left| \frac{F(z) - \Phi(z)}{\Phi(z)} \right| \leq h \end{array} \right\}, \quad h \geq 0 \quad (29)$$

- First 3 conditions specify properties of any cdf.
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- 4th condition states:
fractional error of $\Phi(x)$ is no greater than h .
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- However, h is unknown: $h \geq 0$:
 - **Unbounded horizon of uncertainty.**
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- First 3 conditions specify properties of any cdf.
- 4th condition states:
fractional error of $\Phi(x)$ is no greater than h .
- However, h is unknown: $h \geq 0$:
 - **Unbounded horizon of uncertainty.**
 - No known worst case.

- **The info-gap model of uncertainty, $\mathcal{U}(h), h \geq 0$ is:**
 - **Unbounded family** of nested sets of possible cdf's.
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- **Two properties** of all info-gap models of uncertainty:

$$\mathcal{U}(0) = \{\Phi(z)\}, \quad (\mathbf{Contraction}) \quad (33)$$

$$h < h' \implies \mathcal{U}(h) \subseteq \mathcal{U}(h'), \quad (\mathbf{nesting}) \quad (34)$$

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 - **Unbounded family** of nested sets of possible cdf's.
 - **Non-probabilistic** quantification of uncertainty.
- **Two properties** of all info-gap models of uncertainty:

$$\mathcal{U}(0) = \{\Phi(z)\}, \quad (\text{Contraction}) \quad (35)$$

$$h < h' \implies \mathcal{U}(h) \subseteq \mathcal{U}(h'), \quad (\text{nesting}) \quad (36)$$

§ **Known estimated** and **unknown true**
levels of significance are:

$$\alpha(\Phi) = 2 \left[1 - \Phi \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (37)$$

$$\alpha(F) = 2 \left[1 - F \left(\frac{|\bar{x}_o - T|}{\sigma/\sqrt{N}} \right) \right] \quad (38)$$

§ A common requirement:

- We are concerned that **the estimate, $\alpha(\Phi)$, is overly optimistic:** an under-estimate of the unknown true level of significance, $\alpha(F)$.
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- We **don't know** how much $\Phi(z)$ errs.
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- **We can accept $\alpha(\Phi)$ if it errs no more than ε :**

$$\alpha(F) - \alpha(\Phi) \leq \varepsilon \quad (42)$$

§ Robustness:

- We **don't know** how much $\Phi(z)$ errs.
- We **do know** how much error we can tolerate: eq.(42)
- The **robustness to uncertainty** in the cdf is the max horizon of uncertainty, h , up to which the error is tolerable ...

§ A common requirement:

- We are concerned that **the estimate, $\alpha(\Phi)$, is overly optimistic**: an under-estimate of the unknown true level of significance, $\alpha(F)$.
- **We can accept $\alpha(\Phi)$ if it errs no more than ε :**

$$\alpha(F) - \alpha(\Phi) \leq \varepsilon \quad (43)$$

§ Robustness:

- We **don't know** how much $\Phi(z)$ errs.
- We **do know** how much error we can tolerate: eq.(43)
- The **robustness to uncertainty** in the cdf is the max horizon of uncertainty, h , up to which the error is tolerable:

$\widehat{h}(\varepsilon) =$ **maximum tolerable uncertainty in $F(z)$.**

$=$ **maximum h such that eq.(43) is satisfied.**

$$= \max \left\{ h : \max_{F \in \mathcal{U}(h)} (\alpha(F) - \alpha(\Phi)) \leq \varepsilon \right\} \quad (44)$$

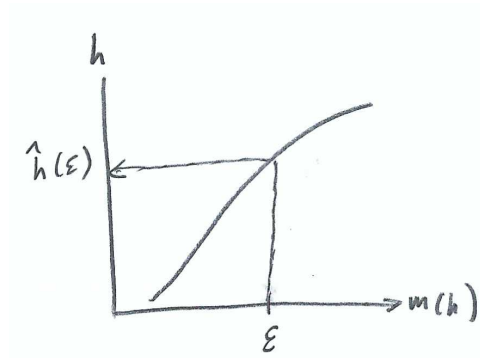


Figure 5: Schematic plot of h vs. $m(h)$.

§ $m(h)$ = inner maximum in definition of robustness:

$$m(h) = \max_{F \in \mathcal{U}(h)} (\alpha(F) - \alpha(\Phi)) \quad (45)$$

This function has two properties (fig. 5):

-

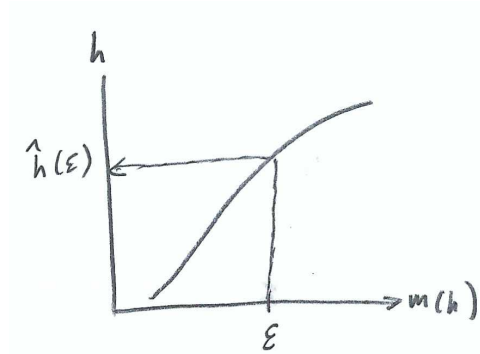


Figure 6: Schematic plot of h vs. $m(h)$.

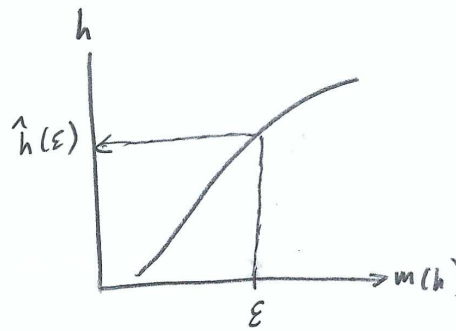
§ $m(h)$ = inner maximum in definition of robustness:

$$m(h) = \max_{F \in \mathcal{U}(h)} (\alpha(F) - \alpha(\Phi)) \quad (46)$$

This function has two properties (fig. 6):

- $m(h)$ increases as h increases.

-

Figure 7: Schematic plot of h vs. $m(h)$.

§ $m(h)$ = inner maximum in definition of robustness:

$$m(h) = \max_{F \in \mathcal{U}(h)} (\alpha(F) - \alpha(\Phi)) \quad (47)$$

This function has two properties (fig. 7):

- $m(h)$ increases as h increases.
- $m(h)$ is the inverse function of $\hat{h}(\varepsilon)$:

a plot of h vs $m(h)$

is the same as

a plot of $\hat{h}(\varepsilon)$ vs ε .

§ Two properties of the robustness:

$$\widehat{h}(\varepsilon) = \frac{\varepsilon}{2\Phi(Q)} \quad (48)$$

§ **Zeroing** in eq.(48):

No robustness ($\widehat{h} = 0$) when requiring no error ($\varepsilon = 0$).

§ **Trade off** in eq.(48):

Robustness increases as greater error is allowed.

1.5 *Second Hypothesis Test: t Distribution*

§ We tested H_0 (p.25) while assuming that σ^2 is **known**.

- What do we do if σ^2 is **unknown**?
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- What do we do if σ^2 is **unknown**?
- We still assume that the **sample is normal**.

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§ We tested H_0 while assuming that σ^2 is **known**.

- What do we do if σ^2 is **unknown**?
- We still assume that the **sample is normal**.

§ Two cases:

- N large.
- N small.

§ Case 1: N large.

- Sample variance:

$$s^2 = \frac{1}{N-1} \sum_{i=1}^N (x_i - \bar{x})^2 \quad (49)$$

-

§ Case 1: N large.

- Sample variance:

$$s^2 = \frac{1}{N-1} \sum_{i=1}^N (x_i - \bar{x})^2 \quad (50)$$

- If $N \geq 25$ (rough number), then s^2 is a **good estimate** of the true population variance, σ^2 .
-

§ Case 1: N large.

- Sample variance:

$$s^2 = \frac{1}{N-1} \sum_{i=1}^N (x_i - \bar{x})^2 \quad (51)$$

- If $N \geq 25$ (rough number), then s^2 is a **good estimate** of the true population variance, σ^2 .
- We can now assume, conditioned on H_0 , that:

$$t = \frac{\bar{x} - T}{s/\sqrt{N}} \sim \mathcal{N}(0, 1) \quad (52)$$

This assumption would be precise if $s^2 = \sigma^2$.

Now we proceed with the test as before.

§ Case 2: N small.

- If $N < 25$ (rough number), then s^2 is **not a good estimate** of σ^2 .
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- The statistic:

$$t = \frac{\bar{x} - T}{s/\sqrt{N}} \quad (53)$$

is broader than $\mathcal{N}(0, 1)$ because \bar{x} and s both display variation.

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- If $N < 25$ (rough number), then s^2 is **not a good estimate** of σ^2 .

- The statistic:

$$t = \frac{\bar{x} - T}{s/\sqrt{N}} \quad (54)$$

is broader than $\mathcal{N}(0, 1)$ because \bar{x} and s both display variation.

- This is a **t statistic** with $N - 1$ degrees of freedom.
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- If $N < 25$ (rough number), then s^2 is **not a good estimate** of σ^2 .

- The statistic:

$$t = \frac{\bar{x} - T}{s/\sqrt{N}} \quad (55)$$

is broader than $\mathcal{N}(0, 1)$ because \bar{x} and s both display variation.

- This is a **t statistic** with $N - 1$ degrees of freedom.
- Recall we assumed the **sample is normal**.

§ We repeat the numerical example, **not knowing** σ^2 .

- Measurements: 1.3, 1.2, 1.4, 1.3, 1.1

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- Measurements: 1.3, 1.2, 1.4, 1.3, 1.1
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- $N = 5$ and $\bar{x}_o = 1.26$.
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§ We repeat the numerical example, **not knowing** σ^2 .

- Measurements: 1.3, 1.2, 1.4, 1.3, 1.1
- Hypothesized temperament: $T = 1.36$.
- $N = 5$ and $\bar{x}_o = 1.26$.
- Sample variance: $s^2 = 0.013 \implies s = 0.1140$.
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§ We repeat the numerical example, **not knowing** σ^2 .

- Measurements: 1.3, 1.2, 1.4, 1.3, 1.1
- Hypothesized temperament: $T = 1.36$.
- $N = 5$ and $\bar{x}_o = 1.26$.
- Sample variance: $s^2 = 0.013 \implies s = 0.1140$.
- The observed t statistic is:

$$t_o = \frac{\bar{x} - T}{s/\sqrt{N}} = \frac{1.26 - 1.36}{0.1140/\sqrt{5}} = -1.961 \quad (56)$$

- The dofs: $5 - 1 = 4$.

- Table of the t distribution:

$\alpha =$	0.1	0.05	0.025	0.01
$\nu = 4 :$	1.533	2.132	2.776	3.747

With 4 dofs:

- The probability of t_4 exceeding **1.533** is **0.1**.
- The probability of t_4 exceeding **2.132** is **0.05**.
- Etc.

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$$\alpha = \mathbf{Prob}(|t| \geq |t_o| \mid H_0) \approx 2 \times 0.07 = 0.14 \quad (57)$$

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- This is **not small, so we cannot reject H_0** .
- The probability of falsely rejecting H_0 is **0.14**.

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Any Questions?